

## Example: Oscillator with Random Phase

Consider the output of a sinusoidal oscillator that has a random phase and amplitude of the form:

$$X(t) = \cos(\Omega_c t + \Theta),$$

where  $\Theta \sim \mathbf{U}([0, 2\pi])$ . Writing out the explicit dependence on the underlying sample space  $\mathbf{S}$  the oscillator output can be written as

$$x(t, \Theta) = \cos(\Omega_c t + \Theta). \quad (31)$$

This random signal falls in the continuous-time, continuous parameter, and continuous amplitude category and is useful in modeling propagation phenomena such as multi-path fading.

The first order distribution of this process can be found by looking at the distribution of the R.V

$$X_t(\Theta) = \cos(\Theta + \theta_o),$$

where  $\Omega_c t = \theta_o$  is a non random quantity. This can easily be shown via the derivative method shown in class to be of the form:

$$f_X(x) = \frac{1}{\pi\sqrt{1-x^2}}, \quad |x| < 1. \quad (32)$$

Note that this distribution is dependent only on the set of values that the process takes and is independent of the particular sampling instant  $t$  and the constant phase offset  $\theta_o$ .

If the second-order distribution is needed then we use the conditional distribution of  $x(t_2)$  as in :

$$f_{x(t_1), x(t_2)}(x_1, x_2) = f_{x(t_2)}(x_2) f_{x(t_1)|x(t_2)}(x_1|x_2) \quad (33)$$

If the value of  $x(t_2)$  is to be equal to  $x_2$  then we require  $\cos(\Theta + \Omega_c t_2) = x_2$ . This can happen only when :

$$\begin{aligned} \Theta &= \cos^{-1}(x_2) - \Omega_c t_2 \quad \text{or} \\ \Theta &= 2\pi - \cos^{-1}(x_2) - \Omega_c t_2, \end{aligned} \quad (34)$$

where  $0 \leq \cos^{-1}(x_2) \leq \pi$ . All other possible solutions lie outside the desired interval  $[0, 2\pi]$ . Consequently the random process at  $t = t_1$  can only take on the values:

$$\begin{aligned} x(t_1) &= \cos(\Omega_c t_1 + \cos^{-1}(x_2) - \Omega_c t_2) \quad \text{or} \\ x(t_1) &= \cos(\Omega_c t_1 - \cos^{-1}(x_2) - \Omega_c t_2) \end{aligned} \quad (35)$$

Thus the conditional distribution of  $x(t_1)$  given that  $x(t_2) = x_2$  is of the form:

$$\begin{aligned} f_{x(t_1)|x(t_2)}(x_1|x_2) &= \left(\frac{1}{2}\right) \delta\left(x_1 - \cos\left[\Omega_c t_1 + \cos^{-1}(x_2) - \Omega_c t_2\right]\right) \\ &+ \left(\frac{1}{2}\right) \delta\left(x_1 - \cos\left[\Omega_c t_1 - \cos^{-1}(x_2) - \Omega_c t_2\right]\right). \end{aligned} \quad (36)$$

Combining Eq. (32) and Eq. (36) we have:

$$\begin{aligned} f_{x(t_1), x(t_2)}(x_1, x_2) &= \left\{ \frac{1}{2\pi\sqrt{1-x_2^2}} \right\} \delta\left(x_1 - \cos\left[\Omega_c t_1 + \cos^{-1}(x_2) - \Omega_c t_2\right]\right) \\ &+ \left\{ \frac{1}{2\pi\sqrt{1-x_2^2}} \right\} \delta\left(x_1 - \cos\left[\Omega_c t_1 - \cos^{-1}(x_2) - \Omega_c t_2\right]\right). \end{aligned} \quad (37)$$

Note here that the second-order PDF depends only on the difference variable  $\tau = t_1 - t_2$ . Let us look at the first-order and second-order moments of the random process  $X(t)$ . The mean of the process is obtained by taking the expectation operator with respect to the random parameter  $\Theta$  on both sides of Eq. (31) keeping in mind that the expectation integral is a linear operation:

$$\begin{aligned}\mu_X &= E_{\Theta}(x_t(\Theta)) = E_{\Theta}[\cos(\Omega_c t + \Theta)] \\ &= E_{\Theta}[\cos(\Omega_c t) \cos(\Theta) - \sin(\Omega_c t) \sin(\Theta)] \\ &= E_{\Theta}[\cos(\Theta)] \cos(\Omega_c t) - E_{\Theta}[\sin(\Theta)] \sin(\Omega_c t).\end{aligned}\quad (38)$$

Since the random parameter  $\Theta$  is uniformly distributed, the above expression reduces to:

$$\mu_X = \cos(\Omega_c t) \left(\frac{1}{2\pi}\right) \int_0^{2\pi} \cos(\theta) d\theta - \sin(\Omega_c t) \left(\frac{1}{2\pi}\right) \int_0^{2\pi} \sin(\theta) d\theta = 0. \quad (39)$$

The variance of the random process  $X(t)$  is obtained by via

$$\sigma_X^2 = E_{\Theta}[(x_t(\Theta) - \mu_X)^2] = E_{\Theta}([x_t(\Theta)]^2) - \mu_X^2 \quad (40)$$

Substituting the mean of the process in the above expression we have:

$$\sigma_X^2 = \left(\frac{1}{2\pi}\right) \int_0^{2\pi} \cos^2(\Omega_c t + \theta) d\theta = \left(\frac{1}{2\pi}\right) \int_0^{2\pi} \left[\frac{1 + \cos(2\Omega_c t + 2\theta)}{2}\right] d\theta = \frac{1}{2} \quad (41)$$

This means that the average power of the random sinusoidal signal  $X(t)$  is

$$P_{\text{ave}}^X = \sigma_X^2 = \frac{1}{2}.$$

Note that this is the same as the average power of a sinusoid where the phase is not random. Let us look at the statistics from the second-order distribution. The correlation between the R.Vs  $x(t_1)$  and  $x(t_2)$  denoted as  $R_{XX}(t_1, t_2)$  is obtained via:

$$\begin{aligned}R_{XX}(t_1, t_2) &= E_{\Theta}[x(t_1)x(t_2)] = \int_0^{2\pi} \cos[\Omega_c t_1 + \theta] \cos[\Omega_c t_2 + \theta] d\theta \\ &= \left(\frac{1}{4\pi}\right) \int_0^{2\pi} \cos[\Omega_c(t_1 + t_2) + 2\theta] d\theta + \left(\frac{1}{4\pi}\right) \int_0^{2\pi} \cos[\Omega_c(t_1 - t_2)] d\theta \\ &= \left(\frac{1}{2}\right) \cos[\Omega_c(t_1 - t_2)].\end{aligned}\quad (42)$$

The covariance of R.Vs  $X(t_1)$  and  $X(t_2)$  denoted  $C_{XX}(t_1, t_2)$  is given by:

$$C_{XX}(t_1, t_2) = R_{xx}(t_1, t_2) - \mu_X(t_1)\mu_X(t_2) = \left(\frac{1}{2}\right) \cos[\Omega_c(t_1 - t_2)]. \quad (43)$$

The correlation coefficient of the R.Vs  $X(t_1)$  and  $X(t_2)$  denoted  $\rho_{XX}(t_1, t_2)$  is:

$$\rho_{XX}(t_1, t_2) = \cos[\Omega_c(t_1 - t_2)]. \quad (44)$$

Looking at the mean and the variance of the random process  $X(t)$  we can see that they are shift-invariant and consequently the process is first-order stationary. The ACF and other second-order statistics of the process are dependent only on the variable  $\tau = t_1 - t_2$ . The

random process  $X(t)$  is therefore a WSS process also. The ACF can then expressed in terms of the variable  $\tau = t_1 - t_2$  as:

$$R_{XX}(\tau) = \left(\frac{1}{2}\right) \cos(\Omega_c \tau). \quad (45)$$

Let us now look at time averages of a single sample function or realization of the random process  $X(t)$ . The sample mean of the random process irrespective of the sample realization that we choose is:

$$\langle \mu_X \rangle_T = \frac{1}{T} \int_{-\frac{T}{2}}^{\frac{T}{2}} \cos[\Omega_c t + \Theta] dt. \quad (46)$$

As  $T \rightarrow \infty$  we have:

$$\lim_{T \rightarrow \infty} \langle \mu_X \rangle_T = 0. \quad (47)$$

The sample mean of the process is therefore independent of the particular ensemble waveform used to calculate the time-average, i.e., independent of the value of  $\Theta$  for the realization. Consequently we have:

$$\begin{aligned} \lim_{T \rightarrow \infty} E \{ \langle \mu_X \rangle_T \} &= \mu_X(t) = 0 \\ \lim_{T \rightarrow \infty} \text{Var} \{ \langle \mu_X \rangle_T \} &= 0. \end{aligned} \quad (48)$$

The random process  $X(t)$  is therefore ergodic in the mean (first-order ergodic). Let us now look at the sample ACF of the random process  $X(t)$ . The sample ACF is again independent of the particular realization of the process as evident from :

$$\begin{aligned} \lim_{T \rightarrow \infty} \langle R_{XX}(\tau) \rangle_T &= \lim_{T \rightarrow \infty} \frac{1}{T} \int_{-\frac{T}{2}}^{\frac{T}{2}} \cos[\Omega_c t + \Theta] \cos[\Omega_c(t - \tau) + \Theta] dt \\ \lim_{T \rightarrow \infty} \langle R_{XX}(\tau) \rangle_T &= \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-\frac{T}{2}}^{\frac{T}{2}} \cos[2\Omega_c t - \Omega_c \tau + 2\Theta] dt + \lim_{T \rightarrow \infty} \frac{1}{2T} \int_{-\frac{T}{2}}^{\frac{T}{2}} \cos(\Omega_c \tau) dt \\ \lim_{T \rightarrow \infty} \langle R_{XX}(\tau) \rangle_T &= \left(\frac{1}{2}\right) \cos(\Omega_c \tau) = R_{XX}(\tau). \end{aligned} \quad (49)$$

The random process  $X(t)$  is therefore ergodic in the ACF (second-order ergodic).

The power-spectrum of this random signal, i.e., the Fourier transform of the ensemble ACF can then be computed as :

$$P_{XX}(\Omega) = \frac{\pi}{2} [\delta(\Omega + \Omega_c) + [\delta(\Omega - \Omega_c)]]. \quad (50)$$

Note that this expression for the power spectrum is identical to the expression for the spectrum of a deterministic sinusoidal signal.